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Delbianco, Fernando
Fioriti, Andrés
González, Germán

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Delbianco, Fernando ¹

Fioriti, Andrés²

González, Germán³

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¹ Professor, Universidad Nacional del Sur– CONICET, fernando.delbianco@uns.edu.ar, San Andrés 800 (8000), Tel: +54291 4595138, Bahía Blanca, Argentina.

² Researcher, Universidad Nacional del Sur – CONICET, andres.fioriti@uns.edu.ar, San Andrés 800 (8000), Tel: +54291 4595138, Bahía Blanca, Argentina.

³ Professor, Instituto de Investigaciones Económicas y Sociales del Sur – CONICET, ghgonza@uns.edu.ar, San Andrés 800 (8000), Tel: +54291 4595138, Bahía Blanca, Argentina.

Abstract:

We propose a novel approach to understand the industrialization pattern of MERCOSUR countries during the last 60 years. We perform an index to measure the geographical bias of manufactured exports between the region and the world and show that regional trade agreements allowed MERCOSUR countries to decrease the bias and increase competitiveness. However, we show that in recent years the bias deepened in favor of exporting to the region a higher proportion of manufactured goods. Our main result is that these economies present a U-shape geographical bias in manufactured exports associated with a bimodal distribution of breaks. Furthermore, we observe that the 1980 and 2000s were the most relevant periods for defining the region trends.

JEL: F1, F10, F15, F14, O14, O54

Keywords: Economic integration, MERCOSUR, Industrialization, Intra-industry trade, exports

1. Introduction

Over the years, there has been a huge deal of papers trying to understand the impact of the regional trade agreements on their members. Since those have been present throughout the whole world and are becoming increasingly popular, in spite of the recent turmoil in the European Union and the apparently cyclical instability of MERCOSUR, it is important to understand whether the impact of this regional trade agreement (RTA) was positive for involucrate semi-industrialized countries or not (Syrquin and Chenery, 1989).

Ceteris paribus, the exploitation of economies of scale and agglomeration increases firms' productivity (Krugman, 1980, Helpman and Krugman, 1989). This reduction in the average costs represents an improvement in the sectoral competitiveness (Harberger, 1998; González, 2006) from non-spurious efforts (Fajnzylber, 1988). Moreover, Puga and Venables (1998) argued that industrialization of a region comes as the expected result

because of market enlargement, in the context of a South-South RTA between partners with similar economies. However, under high asymmetries and slow dismantling of regional trade resistance, the South-South preferential trade liberalization may be, on average, associated with decreased manufacturing production in the smallest country of the agreement (Sanguinetti et al. 2009; Bekerman and Rikap, 2010). From a more traditional perspective, Venables (2003) argued that, when partners have similar productive structures, countries with intermediate comparative advantages between them and the rest of the world benefit from the RTA, at the expense of those partners with extreme or highly concentrated advantages.

Those theoretical branches predict that semi-industrialized countries with large domestic markets experience increasing manufactured exports, principally inter-industry trade, in a South-South RTA context with initial asymmetries. However, if structural and size asymmetries do not exist, diverging patterns between countries may only be transitional because industries may start to disperse as factors of trade resistance are sufficiently reduced. In this last scenario, intra-industry trade dominates within the region. All partners experience an integration of their industries, there is a tendency to intensify their bilateral trade and, as corollary, a tendency to synchronize their business cycles prevails (Frankel and Rose, 1998).

At the end of the process, both cases push extra-regional trade. The difference lies in the evolution of the specialization patterns and the distribution of the benefits between the partners. While regional inter-industry trade promotes the extra-regional exports supported by the static comparative advantages, the regional intra-industry trade pushes manufactured exports towards markets without commercial benefits. For smaller semi-industrialized economies, the first case would show a relative increase in exports of commodities versus manufactured exports, with a prevalence of the flow of those goods to third markets and of the latter to the regional market. In the second case, there would be a relative increase in manufactured exports versus commodities exports, with an increasing participation of third markets in manufactured exports. However, some questions arise: Is it possible that a country experiences both dynamics in a real context of integration and over a sufficiently long period? How could we identify the moment of change and the factors that determine the move from one path to another?

Since our paper is related with a vast literature interested in the evolution of specialization patterns of South American countries (Lo Turco, 2008; Sanguinetti et al., 2009; Bekerman and Dulcich, 2014; among others), we discuss our results in the light of some relevant papers. We focus on the destination of commodities/manufactured exports assuming that these economies have mixed productive structures and that one of its main objectives of entering into a RTA was to use the region as a platform for manufactured exports. Moreover, this target was independent of the initial level of industrialisation, the relative factor endowment and the structure and size of the domestic market. However, this starting point does not imply that we do not consider the influence that those elements may have on the industrialization process. Instead, we verify the consistency of their results by comparing them with our evidence.

Methodologically, our approach is closer to Yeats (1998), Bekerman and Rikap (2010) and Durán Lima and Lo Turco (2010) since they focused on changes on the orientation of trade as a consequence of RTA. We address the dynamic of bilateral, regional, and extra-regional exports of MERCOSUR partners to verify that its economies could not embark in a sustainable and widespread process of industrial competitiveness gains and de-primarization of extra-regional exports. Our main result is that these economies present a U-shape geographical bias in manufactured exports associated with a bimodal distribution of breaks on the data series of their specialization patterns. We used break determination techniques to identify the relevant moments of that U-shape geographical bias. This technique resembles the one used in Delbianco and Fioriti (2014) that attempts to capture the relationship between breaks in the external sector and the occurrence of economic crisis.

The rest of the paper structured as follows: Section 2 explains the data, the construction of the dataset and some descriptive statistics, section 3 presents the breaks testing methodology applied and the results obtained, and section 4 compares the results with the previous literature and discusses some interpretations; finally, section 5 presents the conclusions.

2. Manufactured exports of MERCOSUR countries

Using COMTRADE data about exports of Argentina, Brazil, Paraguay, Uruguay and Venezuela for 1962-2013 period, we obtained the shares of exports to each member, to the region and to the rest of the world (meaning the whole world minus the region) of manufactured and commodities exports.

Although MERCOSUR was born in 1991, González and Delbianco (2020) found that its deployment phase began in 1986 and extended until 1999. While the years prior to 1986 are identified as its prehistory, the 2000-06 phase is considered as an overcoming stage in which MERCOSUR showed its greatest institutional development. However, the authors found that during 2007-15 MERCOSUR returned to a commitment level equivalent to its prehistory. In order to contemplate all these stages and the possibility of different impacts on the specialization process, we consider the period starting in the early experiences of integration of 1960s until 2013, when the five countries in question coincide in MERCOSUR.

Paraguay was suspended between June 2012 and August 2013. Moreover, it was not until December 2013, when the Congress of Paraguay ratified the Protocol of Incorporation of Venezuela into MERCOSUR. However, we incorporated Venezuela into the study because this country began its policy of closer ties with MERCOSUR in February of 1995, being part of MERCOSUR and Andean Group negotiations that conclude with the signing in 2004 of ACE 59⁴.⁵

We adopted the 2-digit level of Standard International Trade Classification (SITC). In order to label each group as manufacture or commodity, we use the usual definition of

⁴ In 2004, MERCOSUR on the one hand and on the other hand the member countries of the Andean Community (CAN) signed the Economic Complementation Agreement N° 59 (ACE 59) by which a reduction of trade barriers between the economies involved is agreed. The Andean Community, previously the Andean Pact or Andean Group, is a trade block comprising the South American countries of Bolivia, Colombia, Ecuador, and Peru, born from the signature of the Cartagena Agreement in 1969. Venezuela was part of it between 1973 and 2006. In August 2017, Venezuela was suspended as its condition of State Party of MERCOSUR through application of the Ushuaia Protocol on Democratic Commitment.

⁵ Bolivia is in the process of accession since 2012 so it does not qualify for inclusion in the study despite also belonging to the Andean community as did it Venezuela.

commodity in trade context -which is a raw material or primary product- for identification⁶. The resulting decomposition is shown in Table I. The distribution results in 28 groups identified as commodities (47%) and 32 groups identified as manufactures (53%).

Table I: SITC-based classification among manufactured (1) and commodity goods (0)

Goods class.	SITC 2-digit	Description	Goods class.	SITC 2-digit	Description
0	0	Live animals	1	53	Dyeing, tanning and colouring materials
0	1	Meat and meat preparations	1	54	Medicinal and pharmaceutical products
0	2	Dairy products and eggs	1	55	Perfume materials, toilet & cleansing preptions
0	3	Fish and fish preparations	1	56	Fertilizers, manufactured
0	4	Cereals and cereal preparations	1	57	Explosives and pyrotechnic products
0	5	Fruit and vegetables	1	58	Plastic materials, etc.
0	6	Sugar, sugar preparations and honey	1	59	Chemical materials and products, nes
0	7	Coffee, tea, cocoa, spices & manufacs. Thereof	1	61	Leather, lthr. Manufs., nes & dressed fur skins
0	8	Feed. Stuff for animals excl. Unmilled cereals	1	62	Rubber manufactures, nes
0	9	Misc.food preparations	1	63	Wood and cork manufs. excluding furniture
1	11	Beverages	1	64	Paper, paperboard and manufactures thereof
0	12	Tobacco and tobacco manufactures	1	65	Textile yarn, fabrics, made up articles, etc.
0	21	Hides, skins and fur skins, undressed	1	66	Non metallic mineral manufactures, nes
0	22	Oil seeds, oil nuts and oil kernels	0	67	Iron and steel
0	23	Crude rubber including synthetic and reclaimed	0	68	Non ferrous metals
0	24	Wood, lumber and cork	1	69	Manufactures of metal, nes
1	25	Pulp and paper	1	71	Machinery, other than electric
0	26	Textile fibres, not manufactured, and waste	1	72	Electrical machinery, apparatus and appliances
0	27	Crude fertilizers and crude minerals, nes	1	73	Transport equipment
0	28	Metalliferous ores and metal scrap	1	81	Sanitary, plumbing, heating and lighting fixt.

⁶ We differ from the usual practice of taking sections 5 to 8 of the SITC as manufactures and the above as primary products. The main argument for this decision is that our classification is based on the final characteristics of the goods, rather than concentrating on the main input used for its production. That is why, vegetable and animal oils, or pulp and paper, turn out to be manufactures, while iron and steel or non-ferrous metals turn out to be commodities.

0	29	Crude animal and vegetable materials, nes	1	82	Furniture
0	32	Coal, coke and briquettes	1	83	Travel goods, handbags and similar articles
0	33	Petroleum and petroleum products	1	84	Clothing
0	34	Gas, natural and manufactured	1	85	Footwear
1	35	Electric energy	1	86	Scientif & control instrum, photogr gds, clocks
0	41	Animal oils and fats	1	89	Misc.manufactured articles, nes
0	42	Fixed vegetable oils and fats	0	93	Special transact. Not class. According to kind
1	43	Animal and vegetable oils and fats, processed	0	94	Animals, nes, incl. Zoo animals, dogs and cats
1	51	Chemical elements and compounds	1	95	Firearms of war and ammunition therefor
1	52	Crude chemicals from coal, petroleum and gas	1	96	Coin, other than gold coin, not legal tender

Source: Author's elaboration

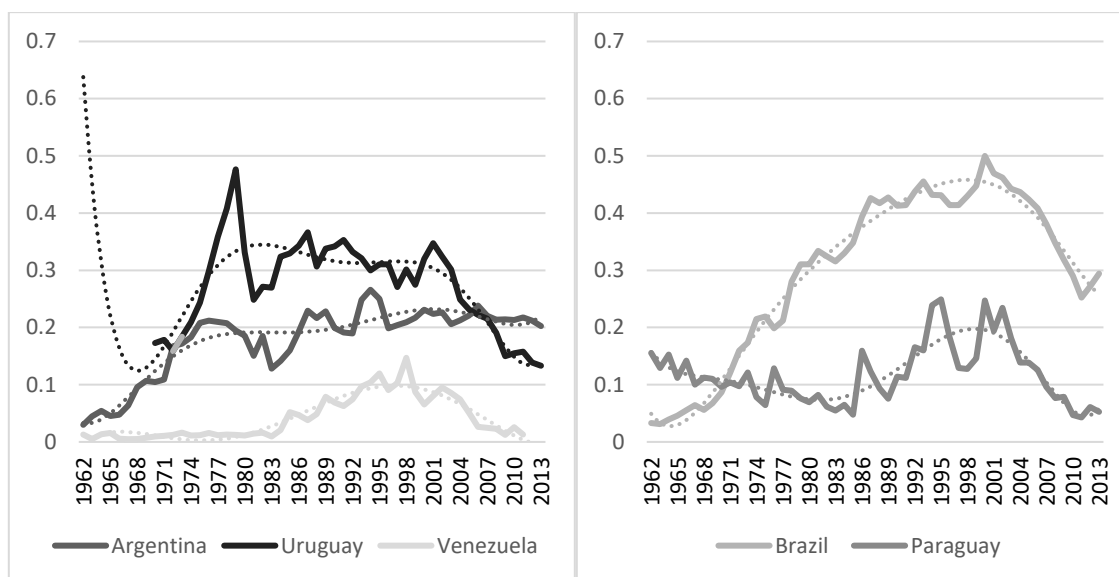
Once all the entries were labeled as manufactured or commodities, we aggregated by country the total amount exported by year. Using the total exports of the year, we obtained the usual ratios:

$$share\ of\ manufactured\ exports_{ij} = \frac{Amount\ of\ manufactured\ exports_{ij}}{Amount\ of\ total\ exports_{ij}} \quad (1)$$

where country i is the exporter and country j is the market. We calculated these ratios for each relevant destination (one for each country of MERCOSUR, one for the whole region, and one for the world minus the region).

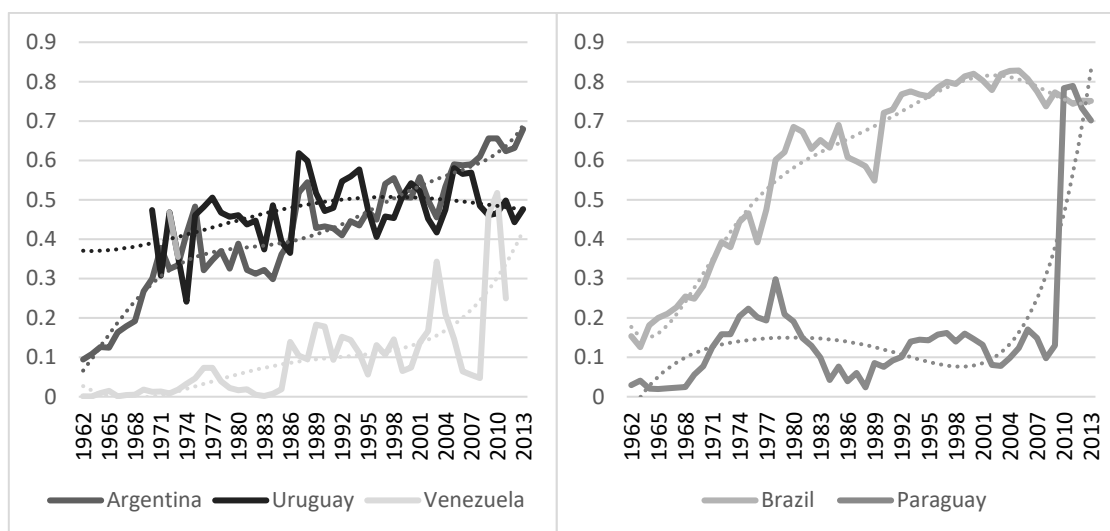
In order to see how these series behaved in the 52-years period considered, we show below the shares of manufactured exports to the world minus the region (Figure I) and to the region (Figure II). The polynomial trend is added to the graphs to obtain a smoother version of the transition over time.

Figure I: Evolution of share of manufactured exports to the rest of the world



Source: Author's elaboration.

Figure II: Evolution of share of manufactured exports to the region



Source: Author's elaboration.

There are several facts that can be observe on the graphs above. At the beginning of the sample, no country of today's MERCOSUR was exporting manufactures for a percentage greater than 20% of total exports to the world, not even to the region. During the following years, each country followed its own path, all of them reaching their maximum values in world market during the 1990s, except for Uruguay that happened previously. Argentina

and Brazil led the de-primarization process based on exporting to the region, but Brazil showed a faster industrialization process and a better trade performance in the world market. Paraguay and Venezuela recently joined a similar path, while Uruguay remained rather stable while looking at the region.

It is not particularly evident on the graph that the RTAs signed within the region fostered the industrialization of the area. We observe a positive trend on the manufactured exports to the region since the 1960s in the five countries considered, while they signed some regional preferential trade and economic complementation agreements between them after the LAIA in the 1980s⁷. In fact, that trend is following a similar phenomenon at a world market level. Finally, in all the cases, an apparent process of primarization of exports to the world market stood out around the year 2000.

Another way of looking at the same issue is to observe the geographic bias of manufactured content of exports⁸ defined as the ratio between the share expressed in (1) for exports to the region and the same share for exports to the world markets. Thus,

$$\text{Geographic bias in manufactured content of exports}_i = \frac{\text{share of manufactured exports to region}_i}{\text{share of manufactured exports to world}_i} \quad (2)$$

where i is the exporter and the shares are calculated by following expression (1) and appropriately assigning j to the region and the world market minus the region.

If this indicator is equal to one it implies geographical neutrality, and this means that firms face exactly the same difficulties and challenges in exporting to the region as to the rest of the world, so the rate of manufactures exported to the region should be the same as the rate of manufactures exported to any market.

⁷ LAFTA was the first proposal for South American economies and signed by Argentina, Brazil, Paraguay and Uruguay in February 1960 and by Venezuela in 1970. However, in view of the impossibility of advancing with a trade-free zone, other Treaty of Montevideo was created LAIA to replace LAFTA in 1980.

⁸ Yeats (1998) used a more disaggregated version of this indicator, and named it regional dimension of the trade specialization process. Volpe Martincus (2003) used a normalized version of this indicator.

Any deviation from the neutrality assumption is reflected with a value different one. Values higher (lower) than one represent a situation in which the ratio between manufactures and commodities exported to the region is greater (lesser) than the same ratio taking the world market as destination. In this scenario, a decreasing (increasing) trend of the bias is the result of a gain (loss) of global competitiveness because industry is exporting relatively more (less) manufactures to markets without access facilities or with greater possibilities of substitution.

Figure III highlights a similar geographical bias behaviour for Argentina, Brazil and Uruguay. All three cases show values above one and follow a U-shape path marking a period of relative expansion of manufactures since the end of 1960s and, after an apparent stagnation during the 1990s, followed by a retreat to initial values from 2000 onwards. In contrast, Paraguay and Venezuela seem to respond to different dynamics than their partners. Although at times they presented some values below one, Venezuela shown a bias in favour of the region over 70% of the years, while Paraguay has done so for 50% of the years. Geographical bias indicator took values above 10 from 2010 to 2013 for Paraguay and above 19 since 2009 to 2011 for Venezuela.

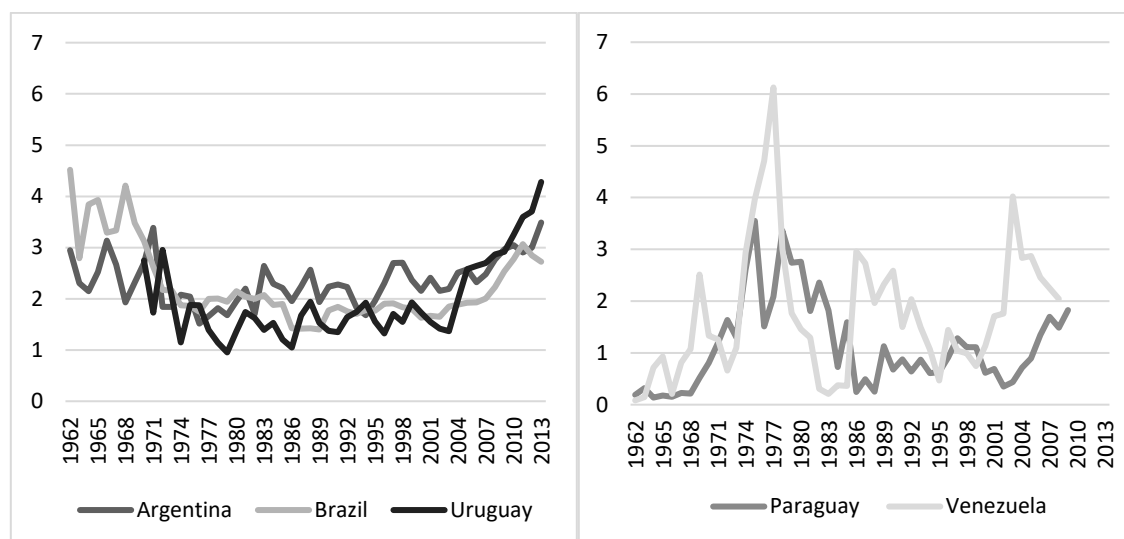
The Paraguayan economy has always concentrated its activities on the production of primary goods and, mainly, agriculture. An important exception to this trend was the period of construction of the Paraguayan-Brazilian Itaipú dam in the 1970s (Richards, 2011) and the significant increase in exports of beef products and soybean oil, with Venezuela being one of its main markets, after 2004 (Swanson, 2012).

The evolution of the Venezuelan bias is associated with the evolution of petroleum exports and its price. Di John (2009) stated that while the period 1920-1965 is characterized by a positive correlation between nonpetroleum sectors growth and the real crude oil price, the subsequent period 1965-1998 is characterized by a significantly negative relationship between these two variables. That evidence prevents a full explanation of industry performance based on the Dutch disease scheme. In any case, this scheme allows us to observe the expansion period of the years 1980s and 1990s in a reasonable way. The devaluations of those years and the depression of the domestic economy helped to boost industrial exports, but when the overvaluation of the bolivar worsened, serious problems of productivity and setbacks in the rate of industrial growth were observed (Hernández and

Villalobos, 1997).

In terms of the peaks observed in the indicator, periods of high growth of crude oil price resulted in higher values for its exports, mainly to third countries. The denominator of the formula (2) is decreasing, and so is the numerator, although to a lesser extent. This combination results in an overestimation of the bias in the manufacturing content of exports to the region. The similarity between the figure of the bias in the manufacturing content of Venezuelan exports and the evolution of the real price of the barrel of crude oil during the period considered is noteworthy⁹.

Figure III: Geographical bias in manufactured exports: region vis-à-vis world market



Note: The graph for Paraguay and Venezuela was cut because values go out of scale after 2010. See text.

Source: Author's elaboration.

It is interesting to note that the contemporaneity of periods of rising values in the partners' indicator –as in the graph for Argentina, Brazil and Uruguay- could represent more regional complementarity in industrial goods and the strengthening of the economic bases for regionalization. However, the other side of the same process is the likely primarization of

⁹ A figure can be seen in this site: <https://www.macrotrends.net/1369/crude-oil-price-history-chart>

exports for global market. This mean that while manufactures were oriented to the partners' economies, the countries maintained their position as suppliers of raw materials and other commodities to the world.

Moreover, these figures show that contrary to what would be expected, the process of intensifying regionalism does not appear to have generated a sharp increment in the geographical bias of manufactured exports. The period of highest growth in intra-regional trade -in particular intra-industry trade- coincides with the period of lower values of the index. Between 1990 and 1999, the four founder members –Argentina, Brazil, Paraguay and Uruguay- significantly increased their Grubel-Lloyd index among them but also with the rest of the world (Kim and Lee, 2003).

During the Commodities Boom, that is the period 2000-2011 -except a sharp down-turn in prices during 2008 and early 2009-, the figures of MERCOSUR's economies –except Brazil- shown an increment of manufactured exports to the region and a rising share of commodity exports –except Argentina-. Graphically, the correspondence between both trends is displayed in figure III as the upward part of the curves.

The results show an overall behaviour consistent with the first scenario presented in the introduction. This means that the integration process helped to sustain the demand for manufactured products in a context of growing globalization. However, the advantage of market access guaranteed by MERCOSUR was not possible to materialize into gains in global competitiveness that would lead to a persistent change in the profile of these countries in world market.

The research presented below has looked to verify the existence of breaks in the specialization paths that give statistical support to the U-shape of the geographical bias of manufactured exports. The identification of the break points on the time series will reduce the difficulty associated with identifying the determinants of that phenomenon.

3. Breaks testing and results

3.1. *Zivot y Andrews (ZA) (1992)*

The classic test of Dickey-Fuller (ADF) with three different specifications (with trend, with constant drift without) without structural changes in the series analyzed generally shows

that the series are non-stationary. However, this result is not surprising when one is working with relatively long time series.

The unit root tests most commonly used as the Dickey-Fuller (1984)¹⁰ or Perron (1989)¹¹ tend not to reject the null hypothesis of a unit root in the presence of structural changes and then generally tend to conclude that the series are not stationary. There are then tests that detect when there was a break in the series analyzed, like Chow, but this first generation requires a priori information about the existence of the break, or use iterations to find such information. A different approach is followed in the work of Zivot and Andrews (1992), which determine endogenously the date of structural change.

The ZA test sequentially analyzes the possible presence of structural changes in the series in each observation, generating dummies in each period. The dummy with a higher level of significance is taken as indicating the period in which the series under study undergoes a change of regime. Removing these sequentially incorporated dummies, the ZA test is then the classic format of a stationarity test (a test of unit root) as the ADF.

ZA is not a structural break test per se. It is, as mentioned, a way to make a stationarity test which tests the null hypothesis of a unit root against the alternative hypothesis of stationarity with a break point at some unknown point in the series.

Therefore, this test has the characteristics of a unit root test with notation similar to Perron but leaving defining the structural break endogenously. The null hypothesis that ZA pose for the three models is:

$$y_t = \mu + y_{t-1} + e_t \quad (3)$$

Thus, it is considered that under H0 the series is integrated series without structural changes. Then, the selection of λ is the result of search for a dummy¹² that achieves a stationary representation of y_t . This means that the alternative hypothesis implies stationarity with a single break. The objective is then to estimate the break (the dummy) that most ponder the alternative of stationarity.

¹⁰ See D. A. Dickey *et al.* (1984).

¹¹ Deeper analysis in the original work, P. Perron (1989).

¹² As in Perron (1989), the breaks are introduced as dummies in any of the three models plausible.

λ is chosen to minimize the one tail t-statistic, testing that $\alpha^i = 1$ for $i = A, B, C$, since small values of the statistic denote rejection of the null. In other terms,

$$t_{\hat{\alpha}^i}[\hat{\lambda}_{\inf}^i] = \inf_{\lambda \in \Lambda} t_{\alpha^i}, \quad i = A, B, C \quad (4)$$

where $\Lambda \in (0,1)$.

As now, the null is specified as in 2.1.1. Then, the equations for return in ZA unit root test are:

$$y_t = \hat{\mu}^A + \hat{\theta}^A DU_t(\hat{\lambda}) + \hat{\beta}^A + \hat{\alpha}^A y_{t-1} + \sum_{j=1}^k \hat{c}_j^A y_{t-j} + \hat{e}_t \quad (5)$$

$$y_t = \hat{\mu}^B + \hat{\gamma}^B DT_t^*(\hat{\lambda}) + \hat{\beta}^B t + \hat{\alpha}^B y_{t-1} + \sum_{j=1}^k \hat{c}_j^B \Delta y_{t-1} + \hat{e}_t \quad (6)$$

$$y_t = \hat{\mu}^C + \hat{\theta}^C DU_t(\hat{\lambda}) + \hat{\gamma}^C DT_t^*(\hat{\lambda}) + \hat{\beta}^C t + \hat{\alpha}^C y_{t-1} + \sum_{j=1}^k \hat{c}_j^C \Delta y_{t-1} + \hat{e}_t \quad (7)$$

where $DU_t(\hat{\lambda}) = 1$ if $t > T\hat{\lambda}$, 0 otherwise; $DT_t^*(\hat{\lambda}) = t - T\hat{\lambda}$ if $t > T\hat{\lambda}$, 0 in any other case.

$\hat{\lambda}$ is the estimated break point after the mentioned procedure.

For each λ , the number of k extra lags, were determined using the Perron's procedure¹³.

Then, the t-statistic is computed. The minimum t obtained on T-2 regressions indicate then the estimated break date.

Once the selection of λ is no longer exogenous through the estimation method, ZA can no longer use the critical values computed by Perron. Now H_0 is rejected if:

$$\inf_{\lambda \in \Lambda} t_{\alpha^i}(\lambda) > k_{\inf \alpha}^i, \quad i = A, B, C \quad (8)$$

with $k_{\inf \alpha}^i$ as the critical size α (toward the left tail of the distribution of the statistic) of the asymptotic distribution of $\inf_{\lambda \in \Lambda} t_{\alpha^i}(\lambda)$. Here, ZA claim that, by definition, these critical values are at least as large as those obtained under an exogenous λ . From this perspective

¹³ This means that the amount of additional regressors k will vary (or may vary at least potentially)

for each selection of λ . (Zivot and Andrews, 1992, p. 255)

then, the unit root test of Perron would be biased to reject the null hypothesis.

3.2. Clemente, Montañes and Reyes (1998)

A particular extension of the test proposed by Perron and Vogelsan (1992) can be found in Clemente et al (1998), for the case of two structural breaks. While the authors recognize that the introduction of a two breaks test was due to Lumsdaine and Papell (1997), this modification was thought for breaks in the trend. Thus, Clemente et al. perform an application to the case of two breaks in the trend.

The null hypothesis of this test is:

$$H0: y_t = y_{t-1} + \delta_1 DTB_{1,t} + \delta_2 DTB_{2,t} + \mu_t \quad (9)$$

while the alternative hypothesis has the following shape:

$$H1: y_t = \mu + d_1 DU_{1,t} + d_2 DU_{2,t} + e_t \quad (10)$$

where $DTB_{i,t}$ is a variable that takes value 1 if $t = TB_i + 1$ ($i = 1,2$) otherwise it is zero, while $DU_{i,t} = 1$ if $t > TB_i$ ($i = 1,2$) and zero otherwise. The periods in time that the breaks take place are denoted as TB_1 and TB_2 , representing the two breaks. The authors assume that the shape of the breaks is $TB_i = \lambda_i T$ ($i = 1,2$) with λ_i between 0 and 1, and $\lambda_2 > \lambda_1$. Moreover, two different specifications are introduced, *additive outliers* (AO) and *innovational outliers* (IO). For the latter, the functional form of the unit root that is tested is

$$y_t = \mu + \rho y_{t-1} + \delta_1 DTB_{1,t} + \delta_2 DTB_{2,t} + d_1 DU_{1,t} + d_2 DU_{2,t} + \sum_{i=1}^k c_i \Delta y_{t-1} + e_t \quad (11)$$

The method used for the test is by minimizing the value of a pseudo t ratio.¹⁴

For the case of AO, the model that is tested is similar but without the deterministic

¹⁴ The functional form of the ratio can be found in Clemente et al. (1997), pp. 177. The assumptions that the authors impose to the possible values of λ , are the same as in Zivot and Andrews (1992), Perron and Vogelsang (1992) and Lumsdaine and Papell (1997).

component:

$$y_t = \mu + d_1 DU_{1,t} + d_2 DU_{2,t} + \hat{y}_t \quad (12)$$

and the test for ρ is realized under

$$\hat{y}_t = \sum_{i=0}^k \psi_{1,i} DTB_{1,t-i} + \sum_{i=0}^k \psi_{2,i} DTB_{2,t-i} + \rho \hat{y}_{t-1} + \sum_{i=1}^k c_i \Delta y_{t-i} + e_t \quad (13)$$

3.3. Results

Table II presents the results distinguishing by origin, destination and identification technique where World (-R) refers to the exports from the origin to the whole world minus the region and the rest of the legends are straightforward.

Table II. Break years by origin, destination and identification technique

Origin						
Destination	DF	ZA I	ZA T	ZA I+T	Clem IO	Clem AO
Argentina						
Brazil	ns	1976	2005	1976	1981s, 2003s,ns	1984s, 2005s,ns
Paraguay	ns	1981	1970	1981*	1966s,1979s,s	1982s,1988s,ns
Uruguay	ns	1970*	1971**	1979**	1967ns,1973s,ns	1966s,2001s,s
Venezuela	s*	2000*	1993	1976	1974s,1998s,ns	1973ns,2000s,ns
World (-R)	ns	1972	1974	1972	1970s, 1982s, ns	1973s,1981ns,ns
Region	ns	1981	1971	1976*	1983s,2002s,ns	1988s,2006s,ns
Brazil						
Argentina	ns	1977	1997	1990	1975s,1988s,s	1978s,1991s,ns
Paraguay	s**	2005	1976	1983	1966ns,1985s,s	1965s,1984s,ns
Uruguay	ns	1978*	1991	1978*	1968s,1977s,s	1976s,1987s,ns
Venezuela	s**	1992	2005*	1999*	1988s,2006s,ns	1994s,2007s,s
World (-R)	ns	2001	1993	1992	1976s,2006s,ns	1975s,1987s,ns
Region	ns	1977	1981	1977	1976ns,1988s,s	1974s,1991s,ns
Paraguay						
Argentina	ns	1984**	1991**	1986**	1969ns,1993s,ns	1993s,2007s,ns
Brazil	ns	1980	2004	1980	1978s,2008s,s	1980s,2007s,ns

Uruguay	ns	2000	1997	2000	1991s,1998s,s	1990s,1997s,s
Venezuela	s**	1972*	1990**	1971**	1966s,1976s,s	1965s,1984s,ns
World (-R)	ns	1986	2003	1992	1990s,2005s,ns	1991s,2005s,ns
Region	ns	1981	2005	2005	1997ns,2008s,s	1970s,2007s,ns
Uruguay						
Argentina	s**	1981*	2005**	1979*	1973s, 1982ns,s	1972s,1981ns,ns
Brazil	s**	1987**	1989	1987**	1972ns, 1985s,s	1980ns,1984s,ns
Paraguay	ns	1995	2000	1995	1994s,2001s,ns	1996s,2004s,ns
Venezuela	s**	1995**	1978**	1983**	1984ns,1993s,s	1980ns,1992s,s
World (-R)	ns	2004	1978	1980	1974s,2002s,ns	1977s,2005s,ns
Region	s**	1987**	1989**	1987**	1973ns,1985s,s	1972ns,1983s,ns
Venezuela						
Argentina	ns	1994*	2000**	1995*	1982s,1998s,ns	1981s,2003s,ns
Brazil	s**	2003	2001**	1999*	1986s,2005s,s	1986s,2004s,s
Paraguay	s*	1993	1969*	1993	1966s,1991s,s	1965s,1990s,ns
Uruguay	ns	2003*	2001*	1991*	1982s,2001s,ns	1981s,2004s,ns
World (-R)	s**	2004**	1980**	2004**	1988s,2005s,ns	1986ns,2005ns,ns
Region	s**	2003	2002*	1999	1985s,2005s,ns	1985s,2004s,ns

Source: Author's elaboration.

s and ns: significative and non-significative. * and **: 5 and 1% level of significance.

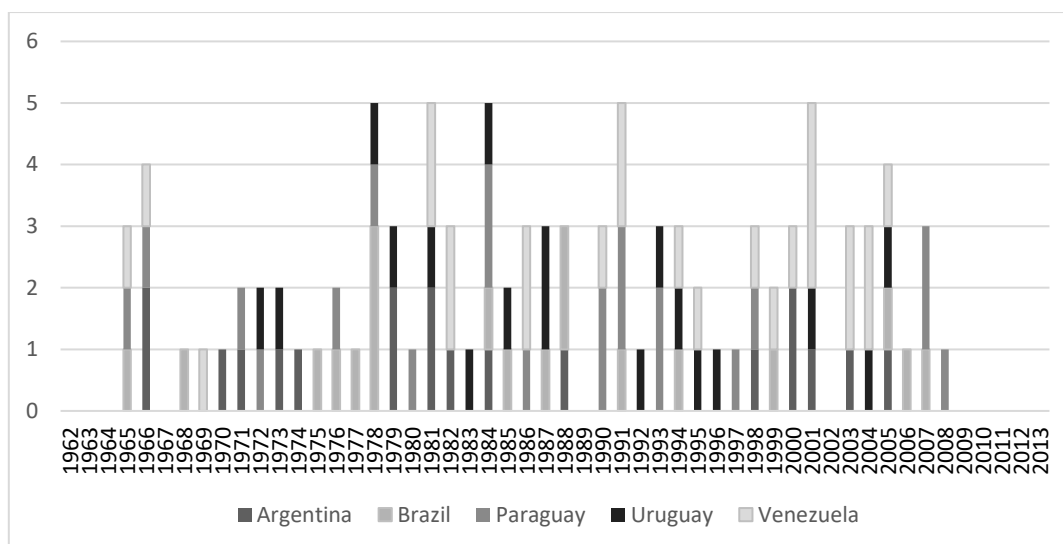
(-R): minus Region; DF: Dickey-Fuller; ZA: Zivot-Andrews; Clem: Clemente; IO and AO: Innovative and Additive Outliers. Exports flows between Paraguay and Venezuela have several missing that generates breaks, phenomenon that not strictly relates to reality (These breaks are highlighted on italics)

There are several facts that we observed in the table II and mainly in the figures IV and V. In the one hand, only three years between 1965 and 2008 show no break in the share of manufactured exports among the current MERCOSUR countries (1967, 1989, and 2002). The accumulation of statistically significant breaks marks a distribution with greater concentration between the 1980s (19 breaks) and the 1990s (20 breaks). However, it is the first of these decades, the only one that has a year with perfect contemporaneity in breaks in four partners. Just like 1984, only 2005 presents perfect contemporaneity between, at least, four economies.

On the other hand, the periods of higher frequency of breaks in exports to a partner are

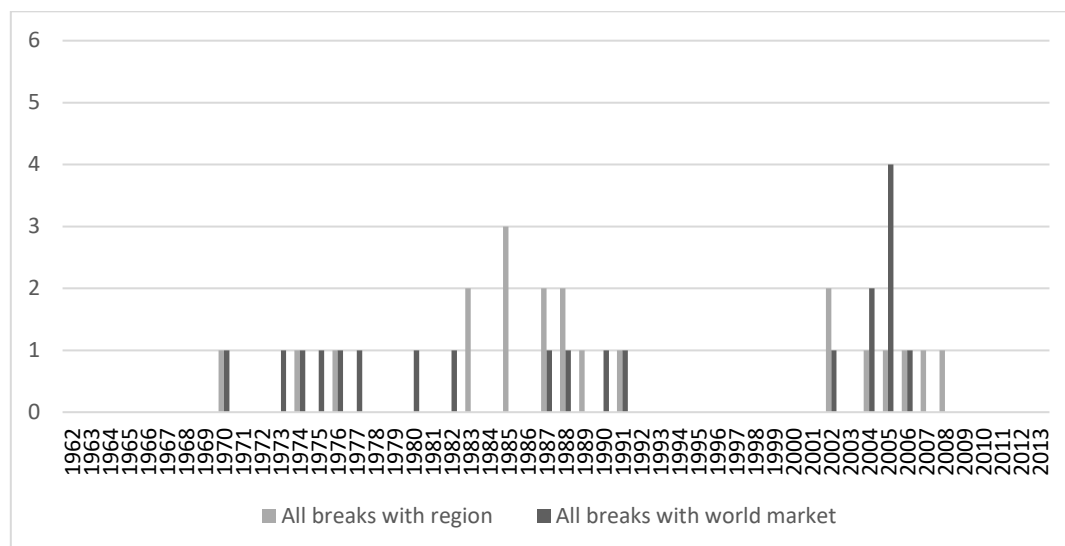
consistent with the periods of higher frequency of breaks in exports to the region in aggregate terms. While this may seem obvious, it is usually the case that a sudden decline in exports to one destination may be offset by higher exports to another destination within the region. Therefore, the data may be showing several breaks in the manufacturing content of exports to different destinations and no change in the manufacturing content of exports to the region. This explains why the number of breaks in figure IV is significantly higher than the number of breaks in the following figure. However, the mid-1980s again, and the mid-2000s appear to be significant moments in the evolution of the specialization pattern of those economies.

Figure IV: Breaks on shares of manufactured exports to partner's market.



Source: Author's elaboration.

Figure V: Accumulation of breaks on shares of manufactured export to the region and to the world market.



Source: Author's elaboration.

Another observation that emerges from figure V is that the distribution of breaks in the manufacturing content of exports to the region is similar to that observed in the same series to the rest of the world. Although there is no perfect contemporaneity, the figure shows greatest concentration in the mid-1980s and in the mid-2000s. In 2005, three MERCOSUR partners experienced breaks in their shares of manufactured exports to the world. Finally, the previous observations describe a series of cumulative breaks with a bimodal distribution whose peaks in mid-1980s and mid-2000s match with the beginning and the end, respectively, of the lower segment of the U-shape series of geographical bias of manufactured exports.

4. Discussion

Both peaks of the series of cumulative breaks coincide with two critical moments of MERCOSUR, according to the k-means periodization proposed in Gonzalez and Delbianco (2019). The first one closes the MERCOSUR's prehistory and gives place to the deployment phase (1986-1999), and the second one opens a regressive transformation process (2007-2015). Cordero and Gonzalez (2018) show, based on the Kojima index of trade intensity, that while the deployment phase identifies a favorable trend to regional

trade in detriment of extra-regional trade, all the partners reversed that trend beginning within a time window that runs from 2000 to 2004.

The “flat” segment of the curve of geographical bias in manufactured exports, whose limits are given by the peaks in the series of cumulative breaks, is consistent with a contemporary increase in intra-industry trade between the MERCOSUR partners and between them and the rest of the world. Kim and Lee (2003) claimed that intra-industrial trade increased drastically after the establishment of MERCOSUR, but this result does not cover only regional trade but also world trade between 1992 and 1999. For his part, Fullerton et al. (2011) remarked the high 2003 values of the Grubel-Lloyd index for Argentina and Brazilian trade with world market, and shown values close to the world average for Uruguay and Venezuela.

In the case of Argentina, de Cicco et al. (2014) verify the correspondence between global and regional Grubel-Lloyd index for the period 1992-2007. However, they showed that the global index was relatively stable between 1997 and 2007 at 35% of its total trade, while it has increasing values over 45% for regional's one, driven mainly by the Argentina-Brazil integration of the automotive sector.

Early on, Yeats (1998) noted that, between 1988 and 1994, MERCOSUR became less internationally competitive in products where trade was reorienting most rapidly toward the region. The Balassa's revealed comparative advantage index for MERCOSUR countries declined for the products that was growing fastest in regional trade. However, more than twenty years later, our evidence implies a positive evaluation of the regionalization process where these semi-industrialized countries develop technological capacities and economies of scale, describing an emerging change of their export profile.

Bekerman and Rikap (2010) argued that the results achieved by Yeats contradict the late reality. MERCOSUR allowed, through learning processes, increase trade with countries outside the block, mainly for Brazil and not so much for the other members. Kim and Lee (2003) provided evidence in accordance with this interpretation. They found that products with high Grubel-Lloyd index values for Argentina and Brazil were mostly in the sectors of chemicals, manufactured goods, and machines and transportation equipment, which are sectors with economies of scale. Fullerton et al. (2011) achieved the same results for the year 2003. However, this is not the case of Paraguay, whose principal sectors were food

and live animals and crude materials, according to Kim and Lee, or beverages and tobacco, and animal and vegetable oils, according to Fullerton. In 2003, Uruguay showed relative high values of the index in beverages and tobacco, and chemical and related products, while Venezuela did it in beverages and tobacco, chemical products and manufactured goods classified by materials.

While Lo Turco (2008) argued that MERCOSUR became a very important destination for exports of medium- and low-technology industrial goods in the period 1985-1999, so grew in relevance natural resources-based goods. Therefore, the structure of regional exports changed slightly in favor of high- and medium-technology industrial goods between 1985 and 2004. This author shows, using a regional variant of Balassa's index, that Brazil maintained its leadership as exporter of industrial goods throughout the period 1985-2004, being the only one of the partners with positive indicator values. Argentina, on the other hand, improved its position in high technology industrial goods but showing regional comparative disadvantages. Similarly, Paraguay revealed a loss of comparative advantage in natural resource-based industrial goods and an improvement in medium and low-technology industrial goods. Uruguay, instead, improved its role as a supplier of natural resource-based manufactures, but also improved its profile in high- and medium-technology manufactures.

Lo Turco found that there was a polarization of the productive specialization between Brazil and smaller MERCOSUR partners, with only a few exceptions in certain sectors and products, and an intermediate position of specialization in the Argentinean case. On one hand, despite the increasing regional intra-industry trade during this period, regional integration has strengthened the comparative advantages of resource-rich countries, reflected in the industrial production and exports mostly based on the exploitation of natural resources. On the other hand, the same process of integration has favoured the forces of agglomeration by making Brazil a beneficiary by increased specialization in industrial activities that are intensive of skilled labor. Moncarz et al. (2016) reaches a similar conclusion. However, they emphasized that these results do not mean that MERCOSUR may have contributed to Brazil's industrialization objective at the expense of the smallest partners, since these countries have also benefited from access to the far larger Brazilian market.

Bekerman and Rikap (2010) concluded that the ten years after the publication of Yeats's analysis showed that MERCOSUR was beneficial to expand trade with third countries and diversify it in the medium and long term based in an interesting process of capacities development. Lucángeli (2007) presented a similar appreciation after verifying the increasing two-way trade in chemicals between Argentina and Brazil. Although our results do not contradict theirs nor support Yeats, we are able to suggest that there will be a setback at the end of the long period under consideration. In fact, Bekerman and Dulcich (2014) remarks the increasing Argentinean dependence on Brazilian imports -principally motor vehicles, plastic and auto-parts industry-, and their results confirm the specialization patterns described by Lo Turco. Volpe Martincus (2003), for his part, uses the index of regional dimension of the trade specialization process to observe an increasing geographical disparity in trade patterns across sectors. Export orientation towards MERCOSUR relative to the rest of the world has increased in those sectors which are favoured by public policy (motor vehicles) and in those sectors in which there is a regional comparative advantage (tobacco) or less comparative disadvantage (textiles and wearing apparel).

Similar results were presented by Reig-Lorenzi and Puchet-Anyul (2018) and Corbella and Sarmiento (2017) using a systemic measure of integration in an input–output framework. The first paper shows that productive integration has been declining since mid-2000s, and the authors emphasized that such behavior was based on the degree of integration between Brazil and Argentina, while Uruguay followed a similar pattern but maintaining a marginal trade relationship with its bigger partners. The second paper shows the remarkable reduction of productive integration in 2005 with respect to 1995. However, the degree of integration of exports of capital goods and of intermediate parts and components has been increasing since 2005. This means a growing dependence –despite the low depth of productive integration- on the regional market for the exports of these types of goods. Gayá and Michalczewsky (2014) observed that during 2003-2011 there was an expansion phase of South American exports. In addition, regional sales expanded at a higher average rate than did shipments to the rest of the world. In particular, the best performance of intra-zone trade was observed between 2004 and 2008. According to the authors, the prevailing model of international insertion was the one theorized by Venables, which predicts inter-

industrial trade based on comparative advantages with the rest of the world and significant intra-industrial regional trade. This remark was valid even for Brazil -the most industrialized country in the region- but with a trend towards the primarization of its exports. They calculate that during 2003-2011, the average intra-industrial trade between the Argentina-Brazil, Argentina-Uruguay, Brazil-Uruguay and Uruguay-Paraguay exceeded 28% of total trade –all but the last one have a steady or increasing trend. The intra-industrial trade between Argentina and Paraguay, and Brazil and Paraguay averaged 15 and 10%, respectively. On the other hand, Venezuela showed very low percentages as it is expected.

From the information provided by the specialized literature and the evidence generated in our study, it is possible to offer a synthesis that, at a minimum, is a well-supported conjecture. It requires empirical comparison using appropriate techniques but presents reasonable evidence in its favour. Considering the dynamics of intra-industrial, regional and global trade, and the series of cumulative breaks with a bimodal distribution, it is possible to develop a unified interpretation of the U-shape of the geographical bias in the manufactured content of exports.

The 1980s saw the exhaustion of the industrial sector that grew during the previous decades under state protection¹⁵. During the 1990s, the open regionalism and the plan of structural reforms opened a possibility of productive reconversion, but the continuing crises of one's own and one's neighbors did not facilitate this process. Two competing forces diverged during those years: The first one was supported by a productive sector based on static

¹⁵ Moncarz et al. (2016) argued that when the strategy of import substitution started to display signs of exhaustion, the regionalism appear as a new means to accelerate the economic development. However, LAFTA suffered the disability to advance with the liberalization process among members during 1960s and 1970s, and LAIA was hampered by the volatility of the 1980s, both in economic and political terms. Then, the South American countries saw in the open regionalism of the 1990s a new possibility of industrialization through the preferential trade agreements.

comparative advantages, homogeneous products, and reasonably detached from the economic and political context and its volatility. The latter was supported by a sector dependent on the development of new technological capabilities -a necessary condition for winning markets-, benefited from preferential trade agreements and the remnants of the industrial policy of previous decades. While external conditions were favourable, manufactured exports were relevant and the second path succeeded in approaching the first. In the early 2000s, these economies surrendered to the great markets' demand for commodities (Delbianco and Fioriti, 2018, and 2019). The industrial sector could survive due to the domestic market, the MERCOSUR trade policy and some competitive corrections to the real exchange rate.

Bekerman and Rodriguez (2005) argued that the harmonization of macroeconomic (especially exchange rate policies) and microeconomic policies, as well as the creation of supranational institutional instances with suitable levels of enforcement, are essentials to obtain successful processes of integration. When looking at MERCOSUR, none of these conditions are met and several abrupt changes in relative prices have affected such harmonization, giving rise to commercial conflicts, the reduction of regional trade, and pushing up the costs of industrialization by protected regional market (in the sense of Cooper and Massell, 1965). As postulated by the theory, the reduction of asymmetries and the prompt dismantling of trade resistance appear to be essential to generate the conditions for developing intra-industry trade and the competitiveness gains for all partners, and an even distribution of the costs of trade diversion (Moncarz et al., 2016).

5. Concluding remarks

We addressed the problem of industry development during the intensification of regionalism in South America, in particular, MERCOSUR. The research focused on trade performance and particularly on the orientation of the manufactured exports as an indicator of the evolution of the industry according to intra-industrial trade models. The new evidence has been confronted with an abundant literature on the evolution of specialization patterns in the economies covered.

The main result has been the identification of a U-shaped bias in the manufacturing content

of exports that is explained by a trend characterized by the exploitation of economies of scale and competitiveness gains, but it did not find sufficient support to consolidate a process of de-primarization of exports. This description corresponds to the experience of the smaller economies of MERCOSUR, but also for Brazil, that is the most industrialized economy in the region.

Using break tests techniques allows us to support the idea that the conjunction of the volatility of the economies during the 1990s and the commodities boom during 2000s may be useful to explain the setback. The first one would not have allowed the consolidation of a path of improvements in technological capabilities, and the second one strengthened the incentives to sustain natural resource-based production.

Our results and interpretation deserve attention in a future deepening of the analysis because they pose a theoretical problem. The analytical framework used does not consider the possibility of a setback in the industrialization process, it does not do so in a scheme with an absence of asymmetries, and even less so in the scheme with a relatively more developed economy than the partners.

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